

1 Heterogeneity in Sectoral Price Stickiness,
2 Aggregate Dynamics and Monetary Policy Pitfalls
3 with Real Shocks

4 Alessandro Flamini*

5 August, 2013

6 **Abstract**

7 Ample differences in sectoral price stickiness is a widely documented fact.
8 This paper shows that in presence of *real* shocks, heterogeneity in sectoral price
9 stickiness plays a key role in the determination of the aggregate dynamics. The
10 larger price stickiness heterogeneity, the smaller the *persistence* of inflation and
11 the *volatility* of inflation, interest rate and output-gap. Thus, two economies
12 with the same average degree of sectoral price stickiness but unlike variance
13 may behave very differently. In terms of monetary policy, they can require *in-*
14 *terest rate paths* that substantially differ both qualitatively and quantitatively.
15 Generally, with real shocks, disregarding the dispersion in sectoral price stick-
16 iness leads policymakers to overvalue the variation and persistence of inflation
17 and output gap.

18 JEL Classification: E31, E32, E37, E52.

19 Key Words: Price Stickiness Heterogeneity, Real Shocks, New Keynesian
20 model, Persistence, Volatility, Monetary Policy Mistakes.

21 **1 Introduction**

22 This paper relates heterogeneity in sectoral price stickiness to aggregate macroeco-
23 nomic dynamics and monetary policy mistakes in presence of real shocks. Real-world

*Current address: Department of Economics, University of Pavia, Via S. Felice 5, 27100 Pavia, Italy. Email: alessandro.flamini@unipv.it. Tel: +39 0382 986 203 (office), +39 333 46 73 515 (mobile).

economies feature multiple sectors differing in terms of nominal rigidities (Altissimo et al. 2006, Blinder et al. 1998, Bils and Klenov, 2004, Dhyne et al. 2006, Nakamura and Steinsson 2008). Most of the research that embeds nominal rigidities, however, adopts a one-sector set up. This is not an innocuous assumption. Indeed, it has been shown both theoretically and empirically that the relation between heterogeneity in sectoral nominal rigidities and macroeconomic dynamics matters in a significant way with monetary shocks. Pioneering this relation, Carvalho (2006) found that the presence of sectoral differences in price stickiness in a New-Keynesian model leads to larger and more persistent real effects in response to monetary shocks. With sticky wage contracts in a Generalized Taylor Economy, Dixon and Kara (2010a) show that for economies with the same average contract length, monetary shocks are more persistent in presence of longer contracts. Empirically, Imbs et al. (2011) show that a Phillips curve based on sectoral estimates implies policy trade-offs remarkably different from one based on aggregate estimates. Focusing on inflation persistence instead of persistence in real variables, Sheedy (2007) finds that heterogeneity in price stickiness leads to less inflation persistence. On the other hand, Dixon and Kara (2010b) argue that considering the distribution of contract length substantially improves the ability of the model to replicate the inflation persistence found in the data.

This previous literature has related heterogeneity in sectoral nominal rigidities to the persistence of real effects caused by *monetary shock*, and to the model performance at reproducing the inflation persistence in the data. The current paper contributes to this literature by showing that in presence of *real shocks*, neglecting heterogeneity in sectoral price stickiness leads to a wrong assessment of the aggregate dynamics, and consequently, to important monetary policy mistakes.

The analysis shows that with real shocks, accounting for the dispersion in sectoral price stickiness, the *persistence* of aggregate inflation turns out to be smaller. The same result occurs with the *variability* of aggregate inflation, interest rate and output-gap. The investigation suggests that these findings are quantitatively important too. Interestingly, this implies that two economies sharing the mean degree of price stickiness but not the variance may respond to real disturbances very differently and thus deserve specific monetary policies. In particular, they can require paths for expected interest rates that remarkably differ both qualitatively and quantitatively. Generally, when real disturbances hit the economy, disregarding the dispersion in sectoral price stickiness leads policymakers to overvalue the variation and persistence of inflation and the variation of the output gap and the interest rate.

59 We believe that these implications pose serious problems to monetary policy de-
60 cisions. In particular, the use of one-sector model would corrupt the projections
61 of the target variables which are fundamental ingredients for the inflation target-
62 ing operating procedure in use at several central banks (Svensson 2010). Thus, this
63 work questions the prominent use of one-sector models for economic estimations and
64 forecasts, a view shared with Imbs et al. (2011).

65 The intuition for the current findings is the following. Heterogeneity in sectoral
66 price stickiness introduces the relative price of sectoral goods into the picture. This
67 variable, affecting sectoral inflations in opposite ways via the demand channel, acts
68 as a buffering device attenuating the impact of aggregate real shocks on aggregate
69 inflation. As a consequence, the whole transmission mechanism of the shock to the
70 economy dramatically changes. Using a graphical AS-AD framework that exactly
71 embeds expectations reveals that the role played by the relative price consists of
72 contracting the shifts of the AS due to real shocks. As a result, aggregate inflation
73 varies less affecting the rest of the economic dynamics.

74 Sectoral and regional heterogeneity in nominal rigidities have been also related by
75 the previous literature to optimal monetary policy obtaining important results. Aoky
76 (2001) and Benigno (2004), respectively in two-sector and two-region economies, show
77 that focusing the policy response on the sector/region with sticky or stickier price
78 maximizes welfare. A different result is obtained by Kara (2010) with a multiple-
79 sector Generalized Taylor Economy where targeting economy-wide inflation results
80 in almost the same welfare of the optimal monetary policy. With respect to this
81 literature, the current paper is different in that abstracts from optimal monetary
82 policy. What it does is a) focusing on the impact of heterogeneity in sectoral price
83 stickiness on the macroeconomic dynamics driven by real shocks, and b) investigating
84 how the path of the expected interest rate changes neglecting heterogeneity in price
85 stickiness. This is carried out contrasting a *two-sector economy* featuring sectoral
86 asymmetry with a *one-sector economy* featuring a degree of price stickiness equal to
87 the average of the sectoral price stickiness in the two-sector economy.

88 The plan of the paper is as follows. Section 2 presents the model where con-
89 sumption habits are introduced into an otherwise standard two-sector New-Keynesian
90 model drawn on Benigno (2004) and Woodford (2011). It derives the non-linear opti-
91 mal conditions, the log-linearized relations used in the following analysis, and reports
92 the calibration of the structural parameters. Section 3 investigates the relation be-
93 tween sectoral heterogeneity in price stickiness and the dynamics of the economy in

presence of cost-push, technology, and preferences shocks. First the main result is illustrated via impulse response functions. Next, the implications in terms of monetary policy mistakes are discussed. Then, the mechanics is explained. Here, the heterogeneity assumption is related to switches in sectoral demands and to their buffering role on aggregate inflation. After that, it is presented a complementary illustration using an appropriate AD-AS graphical framework for the New-Keynesian model. Finally, the analysis is completed looking at autocorrelations, and standard deviations of the endogenous variables. Although this paper focuses on real shocks, Section 4 also offers a parallel with Carvalho (2006) that considers monetary shocks. This comparison shows that introducing the heterogeneity assumption leads to opposite results with monetary and real shocks which, however, always obtain via a contraction in the shift of the AS following the shock. Section 4 also presents three general monetary policy pitfalls in terms wrong assessment of the persistence and volatility in inflation and real activity when policymakers disregard dispersion in sectoral price stickiness. Concluding remarks are in section 5.

2 The model

The economy is populated by a continuum of unit mass of identical infinite-lived households each seeking to maximize

$$U_t = E_t \sum_{T=t}^{\infty} \beta^{T-t} \left\{ \tilde{u}(C_T - \eta C_{T-1}; \bar{C}_T) - \int_0^1 \tilde{v}[H_T(j)] dj \right\},$$

where β is the intertemporal discount factor, C_t represents all interest-rate-sensitive expenditure including investments and is defined as a CES aggregate

$$C_t \equiv \left[(n_s)^{1/\rho} (C_t^s)^{(\rho-1)/\rho} + (n_m)^{1/\rho} (C_t^m)^{(\rho-1)/\rho} \right]^{\rho/(\rho-1)}, \quad (1)$$

of the goods C_t^s and C_t^m which are produced, respectively, by the s and m-sector, with ρ defining their elasticity of substitution and n_s and n_m ($n_s \equiv 1 - n_m$) denoting the number of goods of sector s and m in C_t , respectively. Each sectoral good is, in turn, a Dixit-Stiglitz aggregate of the continuum of differentiated goods produced in the sector

$$C_t^h \equiv \left[n_h^{-\frac{1}{\theta}} \int_{N_h} (C_t^h(i))^{1-\frac{1}{\theta}} di \right]^{\frac{\theta}{\theta-1}}, \quad h = s, m$$

119 where $\theta > 1$ is the sectoral elasticity of substitution between any two differentiated
 120 goods and $N_s \equiv [0, n_s]$, $N_m \equiv (n_s, 1]$. Period preferences on consumption and labour
 121 are modeled as CRRA functions

$$\tilde{u}(C_t - \eta C_{t-1}; \bar{C}_t) = \bar{C}_t^{\frac{1}{\tilde{\sigma}}} \frac{(C_t - \eta C_{t-1})^{1 - \frac{1}{\tilde{\sigma}}} - 1}{1 - \frac{1}{\tilde{\sigma}}}, \quad (2)$$

122

$$\tilde{v}[H_t(j)] \equiv \frac{H_t^{1+\nu}(j)}{1+\nu}, \quad (3)$$

123 where \bar{C}_t is an exogenous preference shock, $H_t(j)$ is the quantity supplied of labour
 124 of type j , $\tilde{\sigma} > 0$ captures the intertemporal elasticity of substitution in consumption,
 125 $0 \leq \eta < 1$ measures the degree of habit persistence, and $\nu > 0$ is the inverse of the
 126 elasticity of goods production¹.

127 The price index for the minimum cost of a unit of C_t is given by

$$P_t \equiv [n_s (P_t^s)^{1-\rho} + (n_m) (P_t^m)^{1-\rho}]^{1/(1-\rho)}, \quad (4)$$

128 with P^s , P^m denoting, respectively, the Dixit-Stiglitz price index for goods produced
 129 in the s and m sector.

130 Preferences captured by equation (1) imply that the optimal sectoral consumption
 131 levels are given by

$$C_t^h = n_h C_t \left(\frac{P_t^h}{P_t} \right)^{-\rho}, \quad h = s, m. \quad (5)$$

132 Financial markets are assumed to be complete so that at any date all households face
 133 the same budget constraint and consume the same amount. Then, utility maximiza-
 134 tion subject to the budget constraint and the no-Ponzi scheme requirement yields the
 135 condition for optimal consumption

$$\lambda_t = \beta E_t \left\{ \frac{[\tilde{u}_c(C_{t+1} - \eta C_t; \bar{C}_{t+1}) - \beta \eta E_t \tilde{u}_c(C_{t+2} - \eta C_{t+1}; \bar{C}_{t+2})] P_t}{\tilde{u}_c(C_t - \eta C_{t-1}; \bar{C}_t) - \beta \eta E_t \tilde{u}_c(C_{t+1} - \eta C_t; \bar{C}_{t+1})} \frac{P_t}{P_{t+1}} \right\}, \quad (6)$$

136 where $\lambda_t \equiv \frac{1}{1+i_t}$ is the price of a one-period nominal bond. Finally, utility maximiza-
 137 tion requires that the optimal supply of labour of type j is given by

$$\Omega_t(j) = \Psi_t \frac{\tilde{v}_h[H_t(j)]}{[\tilde{u}_c(C_t - \eta C_{t-1}; \bar{C}_t) - \eta \beta E_t \tilde{u}_c(C_{t+1} - \eta C_t; \bar{C}_{t+1})]}, \quad (7)$$

¹It is worth noting that the assumption of habit persistence is not necessary to obtain the results shown below. Nevertheless, this real rigidity captures the gradual hump-shaped response of real spending to various shocks and thus is well accepted in the New-Kenesian literature.

where $\Omega_t(j)$ is the real wage demanded for labour of type j and $\Psi_t \geq 1$ is an exogenous markup factor in the labor market assuming that firms are wage-takers.

Moving to production, each household i is assumed to supply all type of labour and is a monopolistically competitive producer of one differentiated good, either $y^m(i)$ or $y^s(i)$. In this economy any firm i belongs to an industry j which, in turn, belongs either to sector s or m . Furthermore, there is a unit interval continuum of industries indexed by j and in each industry there is a unit interval continuum of good indexed by i so that the total number of goods is one. Since in equilibrium all the firms belonging to an industry will supply the same amount, they will also demand the same amount of labour. As a result the total demand of labour in an industry is equal to demand of labor of any differentiated firm in the industry. Next, we assume industry-specific labor as the only variable input

$$y_t^h(i) = A_t [H_t^h(i)]^{\frac{1}{\phi}}, \quad h = s, m,$$

where A_t is a technology shock, $H_t^h(i)$, is the quantity of labour used by the representative firm i in the h -sector to produce good i , and $\phi > 1$, is the elasticity of sectoral output with respect to hours worked.

In equilibrium, market clearing in the goods market requires

$$Y_t^m = C_t^m, \quad Y_t^s = C_t^s, \quad Y_t = C_t. \quad (8)$$

Hence, combining (2), (6), and (8) we obtain the nonlinear version of the New-Keynesian IS relation. Turning to the producers' pricing behaviour, firms in both sectors fix their prices at random intervals following the Calvo (1983) staggered price model and have the opportunity to change their prices with probability $(1 - \alpha)$. Thus, a producer i in the $h = m, s$ sector that is allowed to set its price in period t chooses its new price for the random period starting in t , \tilde{p}_t^h , to maximize the flow of expected profits:

$$\max_{\tilde{p}_t^h} E_t \sum_{T=t}^{\infty} \alpha^{T-t} \lambda_{t,T} \left\{ \tilde{p}_t^h y_T^h(i) - \left[\frac{y_T^h(i)}{A_T} \right]^{\phi} \Psi_T^h \frac{[y_T^h(j)/A_T]^{\nu\phi} P_T}{\bar{C}_t^{\frac{1}{\sigma}} (C_t - \eta C_{t-1})^{-\frac{1}{\sigma}} - \eta \beta \bar{C}_{t+1}^{\frac{1}{\sigma}} (C_{t+1} - \eta C_t)^{-\frac{1}{\sigma}}} \right\},$$

where $\lambda_{t,T}$ is the stochastic discount factor by which financial markets discount random nominal income in period T . Accounting for firm i demand function in sector h , and considering that the firm's pricing decision cannot change the real wage, the f.o.c. is

$$0 = E_t \sum_{T=t}^{\infty} \alpha^{T-t} \lambda_{t,T} \left\{ C_T \left(\frac{\tilde{p}_t^h}{P_T^h} \right)^{-\theta} \left(\frac{P_T^h}{P_T} \right)^{-\rho} - \theta C_T \left(\frac{\tilde{p}_t^h}{P_T^h} \right)^{-\theta-1} \frac{\tilde{p}_t^h}{P_T^h} \left(\frac{P_T^h}{P_T} \right)^{-\rho} - \right.$$

$$\left. \left[-\phi\theta \left(\frac{C_T}{A_T} \right)^\phi \left(\frac{\tilde{P}_T^h}{P_T^h} \right)^{-\phi\theta-1} \frac{1}{P_T^h} \left(\frac{P_T^h}{P_T} \right)^{-\phi\rho} \right] \frac{\Psi_T^h \left[C_t \left(\frac{P_t^h(j)}{P_t^h} \right)^{-\theta} \left(\frac{P_t^h}{P_t} \right)^{-\rho} \frac{1}{A_T} \right]^{\nu\phi} P_T}{\bar{C}_T^{\frac{1}{\sigma}} (C_T - \eta C_{T-1})^{-\frac{1}{\sigma}} - \eta\beta \bar{C}_{T+1}^{\frac{1}{\sigma}} (C_{T+1} - \eta C_T)^{-\frac{1}{\sigma}}} \right\} \quad (9)$$

2.1 Log-linearized relations

We now log-linearize the equilibrium conditions around the steady state where the variables $(Y_t^m, Y_t^s, Y_t, Q_t, \frac{P_{t+1}}{P_t}, \frac{P_{t+1}^s}{P_t^s}, \frac{P_{t+1}^m}{P_t^m})$ are equal to $(Y^m, Y^s, Y, 1, 1, 1, 1)$ and all the shocks are equal to one. Loglinearizing the Euler equation, account being taken of the market clearing condition, leads to the IS relation

$$y_t = \frac{\eta}{1 + \eta(1 + \beta\eta)} y_{t-1} + \frac{1 + \eta\beta(1 + \eta)}{1 + \eta(1 + \beta\eta)} y_{t+1|t} - \frac{\eta\beta}{1 + \eta(1 + \beta\eta)} y_{t+2|t} - \frac{\tilde{\sigma}(1 - \eta)(1 - \eta\beta)}{(1 + \eta + \beta\eta^2)} (\hat{i}_t - \pi_{t+1|t}) + \frac{1 - \eta}{1 + \eta(1 + \beta\eta)} [\bar{c}_t - (\eta\beta + 1)\bar{c}_{t+1|t} + \eta\beta\bar{c}_{t+2|t}] \quad (10)$$

where $\bar{c}_t \equiv \log \bar{C}_t$ which, relaxing the habit persistence assumption, i.e. $\eta = 0$, boils down to the basic New Keynesian IS curve. Next, loglinearizing the f.o.c. for the firm's problem (9) with respect to sector m and s , we obtain

$$\pi_t^h = \kappa^h [\omega + \varphi(1 + \eta^2\beta)] y_t - \kappa^h \varphi \eta y_{t-1} - \kappa^h \varphi \eta \beta y_{t+1|t} + \kappa^h \zeta_h (\rho\omega + 1) q_t \quad (11)$$

$$- \kappa^h [(1 + \omega) a_t + \varphi(1 - \eta)(\bar{c}_t - \eta\beta\bar{c}_{t+1|t}) - \psi_t] + \beta\pi_{t+1|t}^h, \quad h = s, m,$$

where $\omega \equiv \phi(v + 1) - 1$, $\varphi \equiv \frac{1}{(1 - \eta)\tilde{\sigma}(1 - \eta\beta)}$, $q_t \equiv \log \frac{Q_t}{Q}$, $a_t \equiv \log A_t$, $\psi_t^h \equiv \log \Psi_t$, and

$$\kappa^h \equiv \frac{(1 - \alpha^h)(1 - \alpha^h\beta)}{\alpha^h(1 + \omega_h\theta)}, \quad (12)$$

$$\zeta_h \equiv \begin{cases} n_s > 0, & \text{if } h = m \\ -(1 - n_s) < 0, & \text{if } h = s \end{cases} \quad (13)$$

It is worth noting that introducing heterogeneity in sectoral price stickiness the new variable q_t and the new relation between q_t , π_t^s , and π_t^m enter in the model. The latter is captured by both (11) and the law of motion for q_t given by

$$q_t = q_{t-1} + \pi_t^s - \pi_t^m. \quad (14)$$

At this point two remarks are in order. First, considering (12), the shock elasticity of sectoral inflation in (11) is decreasing in the degree of sectoral price stickiness. This implies a shock filtering device for real shocks which is increasing with stickiness as described in Ascari, Flamini and Rossi (2012). Thus, the shock elasticity in the sector whose prices are stickier is smaller. Second, accounting for ζ_h , the elasticities of sectoral inflations to the relative price q_t have opposite sign, and the sector whose prices are more flexible experiences (in absolute value) the larger elasticity. As it will be explained below, different sectoral shock elasticities and relative price elasticities activate a switching demand mechanism that buffers the impact of the symmetric shock on aggregate inflation.

Turning to the exogenous shocks, they are described by

$$\begin{aligned} a_{t+1} &= \gamma_a a_t + \varepsilon_{t+1}^a, \\ \bar{c}_{t+1} &= \gamma_c \bar{c}_t + \varepsilon_{t+1}^c, \\ \psi_{t+1} &= \gamma_\psi \psi_t + \varepsilon_{t+1}^\psi, \end{aligned}$$

where $E_t(\varepsilon_{t+1}^h) = 0$, $h = a, c, \psi$. Log-linearizing the price index (4) we obtain aggregate inflation

$$\pi_t = n_s \pi_t^s + n_m \pi_t^m, \quad (15)$$

and substituting the sectoral inflations we obtain aggregate inflation in terms of lagged, current, and expected output gap, the relative price, expected inflation, and the exogenous shocks

$$\begin{aligned} \pi_t &= [\omega + \varphi(1 + \eta^2 \beta)] (n_s \kappa^s + n_m \kappa^m) y_t - \varphi \eta (n_s \kappa^s + n_m \kappa^m) y_{t-1} + \beta \pi_{t+1|t} \\ &\quad - \varphi \eta \beta (n_s \kappa^s + n_m \kappa^m) y_{t+1|t} - n_s n_m (\kappa^s - \kappa^m) (\rho \omega + 1) q_t \\ &\quad - (n_s \kappa^s + n_m \kappa^m) [(1 + \omega) a_t - \psi_t] - \varphi (1 - \eta) (n_s \kappa^s + n_m \kappa^m) (\bar{c}_t - \eta \beta \bar{c}_{t+1|t}). \end{aligned} \quad (16)$$

This aggregate New-Keynesian Phillips curve based on sectoral inflations boils down to the standard New-Kenesian Phillips curve with habit persistence when sectoral symmetry is imposed², i.e. $\alpha^s = \alpha^m$. Finally, the model is closed with a Taylor rule

²Assuming also no habit persistence, i.e. $\eta = 0$, we obtain the basic New-Keynesian Phillips curve.

197 describing the behaviour of the central bank

$$i_t = \delta_0 i_{t-1} + (1 - \delta_0) \delta_1 \pi_t + (1 - \delta_0) \delta_2 y_t. \quad (17)$$

198 **2.2 Calibration**

199 Table 1a reports the calibration for the structural parameters based on the previous
200 literature. The degree of habits persistence η is set to 0.7; the elasticity of intertem-
201 poral substitution in consumption is $\tilde{\sigma} = 2/3$; the intertemporal discount factor is
202 $\beta = 0.9975$ (3% per year); the coefficients of the Taylor rule are $\delta_0 = 0.8$; $\delta_1 = 1.5$;
203 $\delta_2 = 0.5/4$. These values configure a quite standard calibration consistent, for exam-
204 ple, with Smets and Wouters (2007). Let us now turn to the remaining parameters.
205 The inverse of the elasticity of goods production (the inverse of Frish elasticity) ν is set
206 to 1.17 as estimated by Fernández-Villaverde, Guerrón-Quintana and Rubio-Ramírez
207 (2010)³. Following Rotemberg and Woodford (1997), the elasticity of sectoral output
208 with respect to hours worked, $1/\phi$, is set to 0.75 and the sectoral elasticity of substi-
209 tution between any two differentiated goods θ is set to 7.88 (average markup $< 15\%$).
210 The elasticity of substitution between C_t^s and C_t^m in the CES consumption aggregate,
211 ρ , is difficult to calibrate without specifying the type of industry the sectors belong
212 to. We then assume that the m and the s sector refer to the manufacturing and ser-
213 vices sectors respectively, and set ρ to 1 (Cobb-Douglas aggregator). Since this work
214 aims to insulate the impact of the asymmetry in sectoral price stickiness on economic
215 dynamics in presence of real shocks, sectors size is set to be equal, i.e. the number of
216 firms in the s-sector is $n_s = 0.5$; and in the m-sector is $n_m = 1 - n_s$. Finally, the AR
217 coefficients of the exogenous processes are $\gamma_a = \gamma_\psi = \gamma_c = 0.95$, for any shock the
218 variance is $\sigma_\varepsilon^2 = 0.009^2$, and the sectoral degree of price stickiness α_h , $h = s, m$ is let
219 free to vary in the range $\{0.5, 0.6, 0.7, 0.8, 0.9\}$ as described in the analysis below.

220 **Robustness** To check for the robustness of the results, we experimented an alter-
221 native calibration which is as close as possible to the one in Carvalho (2006)⁴. Table

³This source for ν is motivated by the fact that these authors use a set of priors nearly identical to the one proposed by Smets and Wouters (2007) in a similar DSGE model but specify the relation between the utility and labor in the same way of the current paper, which differs from Smets and Wouters (2007).

⁴Carvalho (2006) considers two alternative calibrations for ν and θ : either $1/\nu = 0.5$ and $\theta = 11$ or $1/\nu = 1.5$ and $\theta = 5$. Here, for sake of simplicity but without loss in generality, we take the

1*b* only reports the parameters value of this second calibration which differ from the first calibration. Here there are no habits in consumption, i.e. $\eta = 0$, the elasticity of intertemporal substitution in consumption is $\tilde{\sigma} = 1$, the elasticity of sectoral output with respect to hours worked is $1/\phi = 1$ (linear technology), and the inverse of the elasticity of goods production is $\nu = 1$. Finally, the sectoral elasticity of substitution between any two differentiated goods θ and the elasticity of substitution between C_t^s and C_t^m in the CES consumption aggregate ρ are both set equal to 8.

3 Sectoral heterogeneity in price stickiness and economic dynamics

We now focus on the relation between sectoral heterogeneity in price stickiness and the dynamics of the economy in presence of positive cost-push, technology and household preferences shocks. All the shocks are supposed to hit symmetrically both sectors of the economy. The analysis first shows qualitatively via impulse response functions how dispersion in sectoral price stickiness affects the economic dynamics. Next, it assesses the policy mistakes for a central bank that ignores the presence of dispersion in sectoral price stickiness. Then the mechanics is explained focusing on the impact of the asymmetry assumption on the sectoral and aggregate AS. The mechanism at work is also illustrated using an appropriate AD-AS diagram for the New-Keynesian model. Finally, the investigation is completed looking at autocorrelations, and standard deviations of the endogenous variables.

3.1 Impulse responses to real shocks

Figures 1 shows the impulse response functions to a preference shock (Panel a), and to symmetric cost-push (Panel b) and technology shocks (Panel c) in presence of sectoral symmetry and asymmetry in the degree of price stickiness.

[INSERT FIGURE 1 HERE]

average values over these two calibrations, that is $1/\nu = 1$ and $\theta = 8$. It is worth noting that in Carvalho (2006) ρ is set equal to the sectoral elasticity of substitution θ , and it is reported that alternative calibrations for ρ including the case in which the CES is a Cobb-Douglas do not change the substantive findings.

247 The first row of each Panel considers a two-sector economy under the *symmetry*
 248 assumption, i.e. price stickiness is the same in both sectors. Here the probability that
 249 a firm does not have the opportunity to optimize its price in a given period captured by
 250 α_s and α_m is set to 0.7. This value implies that, on average, firms revise their prices
 251 a bit less than every 3.5 quarters, which is consistent with Fernández-Villaverde,
 252 Guerrón-Quintana and Rubio-Ramírez (2010) and seems a natural benchmark of
 253 longer and shorter pricing cycle for firms belonging to different sectors. The second
 254 row, instead, considers the same economy under the *asymmetry* assumption but with
 255 the same mean of the symmetric economy. Specifically, the sectoral price stickiness
 256 are $\alpha_s = 0.9$, $\alpha_m = 0.5$, resulting in the mean of 0.7 over the two sectors⁵. What is
 257 interesting here is that each panel reveals a common behaviour of aggregate inflation
 258 in response to an increasing degree of price stickiness asymmetry. This behaviour can
 259 be described as follows:

260 **Variance of sectoral price stickiness and aggregate inflation.** *For any*
 261 *shock considered, for a given mean value of price stickiness, the larger the variance*
 262 *(i.e. the larger the asymmetry), the lower the deviation of inflation from its steady*
 263 *state value, the lower the initial impact of the shock and the lower the persistence of*
 264 *the response to the shock⁶.*

265 Before explaining the mechanism that generates this result, we describe its im-
 266 plications on the behaviour of the interest rate and then on the economic dynamics.
 267 Starting with the preference shock, Panel a, it is worth noticing that this shock im-
 268 pacts on both the demand and supply side of the economy⁷, and these impacts go
 269 in opposite directions⁸. Plots for y and π then illustrate these effects, positive on
 270 the output gap and negative on inflation. Thus, two contrasting forces act on the

⁵The measure of price stickiness captured by α in the Calvo's (1987) scheme varies a lot in the empirical literature ranging from 0.35 in Christiano et al. (2005) to 0.91 in Smets and Wouters (2003).

⁶Further analysis available upon request shows that this result is independent on the mean and the variance of sectoral price stickiness.

⁷The preference shock is a shock to the utility function so that it affects, on the one hand, the Euler equation and therefore the aggregate demand and, on the other hand, the marginal rate of substitution between labour and consumption entering in the optimal supply of labour, and therefore the aggregate supply.

⁸Indeed in equation (10) the coefficient of the preference shock is positive and in equation (16) it is negative.

interest rate via the Taylor rule. Now, under symmetry, first row, it occurs that the
fall in inflation affects the interest rate more than the increase in the output gap
leading the interest rate to fall in the initial periods. Yet, breaking the symmetry,
second row, inflation is less affected by the shock and, therefore the negative impact
on the interest rate is attenuated. This leaves the interest rate more exposed to the
positive impact of the output gap. Consequently, the monetary policy turns out to
be more active in the subsequent periods in order to focus on the stabilization of the
output gap. Hence, given the relation between the variance of sectoral price stickiness
and aggregate inflation, monetary policy can better stabilize the output gap under
asymmetry.

Moving to a cost-push or technology shock, the economy turns out to be hit only
through the supply side. Thus the larger the asymmetry, the less the shock affects
aggregate inflation. This implies that the interest rate has to respond less via the
Taylor rule and thus that the shock propagates less to the output gap. Accordingly,
Panels b and c show that the policy response in the asymmetric economy tends
to be remarkably milder than in the symmetric one. Similarly, the deviations of
output-gap and inflation from their steady state values in the asymmetric economy
tend to be remarkably smaller than in the symmetric economy. Summing up, with
supply shocks, heterogeneity in sectoral price stickiness tends to buffer exogenous
disturbances requiring a less active monetary policy.

3.2 Monetary policy mistakes ignoring sectoral price stickiness heterogeneity

The previous findings imply that two two-sector economies sharing the same mean
degree of price stickiness but not the same variance respond differently to exogenous
real disturbances and therefore require specific monetary policies. This implication
poses serious problems in terms of monetary policy mistakes for central banks that
neglects sector price stickiness heterogeneity. To illustrate this point, we can consider,
for instance, a two-sector economy where $\alpha_s = 0.9$ and $\alpha_m = 0.5$ so that the average
degree of price stickiness is $\alpha = 0.7$. In this economy, we let the central bank assume
that $\alpha_s = \alpha_m = 0.7$, or, equivalently for what is shown below, only consider the aver-
age price stickiness value disregarding the dispersion across sectors. Then, exogenous
shocks hit and the central bank responds. Yet, the actual response will be different
from the response that would occur if the central bank considered heterogeneity in

304 sectoral price stickiness, i.e. the correct response. To develop this argument, let us
305 define as *monetary policy* the current and expected interest rate decisions, that is
306 the interest rate path following the shock⁹. There are, then, two monetary policies:
307 the actual and the correct one, and we can ask what the policy mistakes are for the
308 central bank that neglects the dispersion in sectoral price stickiness.

309 Figure 2 provides a *prima facie* answer contrasting in each panel the IRFs of the
310 interest rate capturing the actual (dashed) and the correct policy (solid). Here each
311 panel lies on a row referring to a calibration and on a column referring to a shock.

312 [INSERT FIGURE 2 HERE]

313 Figure 2 shows that asymmetry in price stickiness always makes monetary policy
314 remarkably different. With a cost-push shock, ε^ψ , the actual policy is uniformly
315 *tighter* than the correct one, while with a technology shock, ε^a , the actual is uniformly
316 *easier* than the correct; these results holding with both calibrations. Left with a
317 preference shock, ε^c , under the first calibration the actual policy is *less tight* than
318 the correct one, except for the initial three periods. Under the second calibration
319 the actual policy is fully wrong: it is *easy* instead of *tight* as required by the correct
320 policy.

321 Visual inspection thus reveals that neglecting asymmetry in price stickiness leads
322 to quantitative and qualitative policy mistakes. At this point, to sharpen our evaluation
323 of these policy mistakes we need to measure the policy error. For this purpose, let us
324 introduce a policy error statistic in two steps: first computing the distance between
325 the interest rate path under symmetry and asymmetry at each point in time. Then
326 dividing this distance by the value of the policy under asymmetry at each point in
327 time. What results is therefore the *percentage deviation of the actual policy from the*
328 *correct policy*. Figure 3 reports this statistic for the policy errors. As before, each
329 panel lies on a row that refers to a calibration and to a column that refers to a shock.

330 [INSERT FIGURE 3 HERE]

331 In addition, now, each panel shows for every point in time the measure (in per-
332 centage terms) and the type of the error (e.g. policy tighter, easier, etc.). To fix
333 the ideas, we can consider the policy errors with a cost-push shock under the first

⁹This definition is introduced to avoid confusion between the *monetary policy rule*, which here is the Taylor rule and is invariant to the symmetry/asymmetry case, and the *monetary policy*, that is the decisions (current and expected) taken by the central bank using the policy rule.

calibration (row1, column 1). Looking at the tenth period subsequent to the shock, the actual policy turns out to be 100% tighter than the correct one. Now, what Figure 3 reveals is that the policy errors are generally very large, with actual policies that compared to the correct ones can be easier (black), tighter (blue/dark grey), less tight (white) between 40% and 145% over the first ten periods. Or with policies that are easy instead of tight (yellow/light grey), with a distance between the two which is no less than 150% of the (absolute) value of the correct tight policy. Summing up, Table 2 reports that the percentage deviation of the actual policy from the correct policy averaged over the first ten periods is between 50% and 100% for the cost-push and technology shock, and between 60% and 180% for the preference shock.

Now, before deepening the investigation with a quantitative assessment of the implications of this policy mistakes, it is worth stopping for a natural question: what drives these findings?

3.3 Mechanics

When sectors differ in terms of stickiness, sectoral inflations differ after a shock and the stickier sector experiences a smaller change in inflation¹⁰. As a result the relative price between sectors kicks in as shown by equation (14) and tends to divert the demand from the sector whose goods are relatively more expensive to the sector whose goods are relatively cheaper. As expected, in the former sector inflation falls while in the latter increases. Yet, these sectoral inflation changes caused by q , beyond differing in direction, differ also size wise because the change in marginal costs caused by the demand change impacts less inflation in the stickier sector. This is due to the fact that the elasticity of sectoral inflation to the relative price is decreasing in sectoral stickiness as shown in (11) and therefore is smaller in the stickier sector. Such a difference in the impact of the relative price on sectoral inflations implies a fall in aggregate inflation and therefore a buffering role played by the switching demand mechanism. It is worth noting that the sector where prices are more flexible, while exposing aggregate inflation more to the shocks, is also the one through which the switching demand mechanism acts the most to buffer the shock.

The buffering device activated by heterogeneity in price stickiness can be conveniently illustrated with a New Keynesian AS-AD framework that gives special atten-

¹⁰This holds no matter what the shock is: with a supply shock through the shock filtering mechanism and with a demand shock through a different slope of the Phillips curve.

tion to the role played by expectations.

3.3.1 A New Keynesian AS-AD graphical illustration of the role of price stickiness dispersion

The AD-AS diagram is presented in the (π_t, y_t) space. For sake of simplicity but without loss in generality the assumptions of habit persistence and interest rate smoothing are relaxed, i.e. $\eta = 0$ and $\delta_0 = 0$. Drawing the diagram, it is worth recalling that in steady state all the expectations terms are zero.

To obtain the AD we join the Euler equation and the Taylor rule and solve for current inflation as a function of the current output gap given the output gap and inflation expectations

$$\pi_t = - \left(\frac{1 + \tilde{\sigma}\delta_2}{\tilde{\sigma}\delta_1} \right) y_t + \frac{1}{\tilde{\sigma}\delta_1} y_{t+1|t} + \frac{1}{\delta_1} \pi_{t+1|t}. \quad (18)$$

Solving this equation forward leads to

$$\pi_t = - \left(\frac{1 + \tilde{\sigma}\delta_2}{\tilde{\sigma}\delta_1} \right) y_t + \frac{\delta_1 - 1 - \tilde{\sigma}\delta_2}{\tilde{\sigma}\delta_1} \sum_{\tau=1}^{\infty} \delta_1^{-\tau} y_{t+\tau|t}. \quad (19)$$

Equation (19) shows that the relation between current inflation and output gap shifts in the (π_t, y_t) space according to the expected path of the output gap. Thus, if breaking the symmetry changes this expected path there will be a shift of the AD with respect to its position under symmetry.

Regarding the AS, equation (16) can be solved forward to obtain

$$\begin{aligned} \pi_t = & (\omega + \varphi) \frac{1}{2} (\kappa_s + \kappa_m) y_t + \frac{1}{2} (\omega + \varphi) (\kappa_s + \kappa_m) \sum_{\tau=1}^{\infty} \beta^\tau y_{t+\tau|t} \\ & - \frac{1}{4} (\rho\omega + 1) (\kappa_s - \kappa_m) \sum_{\tau=0}^{\infty} \beta^\tau q_{t+\tau|t} + \frac{1}{2} (\kappa_s + \kappa_m) \sum_{\tau=0}^{\infty} \beta^\tau \psi_{t+\tau|t}, \end{aligned} \quad (20)$$

which under symmetry boils down to

$$\pi_t = (\omega + \varphi) \frac{1}{2} \kappa y_t + \frac{1}{2} (\omega + \varphi) \kappa \sum_{\tau=1}^{\infty} \beta^\tau y_{t+\tau|t} + \frac{1}{2} \kappa \sum_{\tau=0}^{\infty} \beta^\tau \psi_{t+\tau|t}, \quad (21)$$

where $\kappa = \kappa_s = \kappa_m$. For a given expected path of the cost-push shock, equation (20) shows that current inflation depends on the expected paths of the output gap and the relative price, while equation (21) shows that it depends only on the expected path of the output gap.

3.3.2 Cost-push shock case

Figure 4 reports the *steady state equilibrium A* and the *off-steady state equilibria* in presence of a cost-push shock under symmetry *B* ($\alpha_m = \alpha_s = 0.7$), and asymmetry *C* ($\alpha_m = 0.6, \alpha_s = 0.8$) and *D* ($\alpha_m = 0.5, \alpha_s = 0.9$).

[INSERT FIGURE 4 HERE]

Off the steady state, Figure 4 takes a snapshot of π_t and y_t in the initial period in which the shock hits the economy. It does not leaves aside the subsequent periods though. Indeed, expectations on the economic dynamics determine the position of the AD and AS (not their slopes) as shown by equations (19-21). Describing the Figure, the red/grey and the black color refer to symmetry and asymmetry respectively, while the solid lines refer to the steady state equilibrium, and the dash and dots lines refer to off-steady state equilibria. It is worth noting that in steady state the AD (downward sloped) under symmetry and asymmetry coincide, while the AS (upward sloped) is steeper under asymmetry¹¹.

Figure 4 shows that under symmetry, when the shock hits the economy the equilibrium is no longer A but B. What happens is that with the shock the expected path for inflation becomes positive so that, via the Taylor rule, the expected path for the output gap becomes negative. As shown by equations (21 and 19), this leads the AD to shift left and the AS to shift up (the impact of the shock prevails on the deflationary path for the output gap). Breaking the symmetry the economy moves either to *C* (low asymmetry) or to *D* (high asymmetry). With respect to equilibrium B, equilibria *C* and *D* feature a smaller increase in inflation and a smaller fall in output. Furthermore, the larger the asymmetry, the larger the departure from *B* and the larger these effects. Indeed, regarding the AS, for a given current value and expected path of the shock, under asymmetry the shift is determined by the expected path of the output gap and the relative price, summarized in (20) by the terms $\sum_{\tau=1}^{\infty} \beta^{\tau} y_{t+\tau|t}$

and $\sum_{\tau=0}^{\infty} \beta^{\tau} q_{t+\tau|t}$ respectively. In contrast, given the shock, under symmetry the shift is determined only by the expected path of the output gap. Thus, in buffering the shock the *deflating mechanism* provided by the output gap is supported by the *switching*

¹¹Breaking the symmetry, the elasticity of inflation to the output gap and the shock increases. This effect, visible in the AS steeper under asymmetry and due to the convexity of the relation between κ and α , exerts a minor impact on the macroeconomic dynamics.

414 *demand mechanism* provided by the relative price. This is shown by the *negative*
 415 expected path for q reported in Panel b of Figure 1 associated with the positive elas-
 416 ticity of aggregate inflation to the relative price in equation (20). Thus the expected
 417 path of the relative price buffers the shock in addition to the expected path for the
 418 output gap in (20). As a result of the presence of the switching demand mechanism,
 419 the AS shifts up less under asymmetry.

420 Regarding the AD, the smaller shift to the left under asymmetry is always due to
 421 the presence of the relative price. Indeed, the smaller impact of the shock on inflation
 422 due to the switching demand mechanism implies, via the policy response, a smaller
 423 deflating role imposed to the output gap, which is captured by a smaller value of
 424 the term $\sum_{\tau=1}^{\infty} \beta^{\tau} y_{t+\tau|t}$ in (19). Summing up, breaking the symmetry results in different
 425 impacts of the cost-push shock on sectoral inflations which, activating the relative
 426 price, allows a partial return of both AS and AD to their pre-shock positions. Hence
 427 breaking the symmetry results in lower output and inflation volatility.

428 **3.4 Persistence and volatility of the macrovariables vs sec-** 429 **toral price stickiness heterogeneity**

430 The Impulse Response analysis has signalled two consequences of the presence of
 431 sectoral price stickiness asymmetry in the working of the economy. On the one
 432 hand it affects the macroeconomic dynamics. On the other hand, if it is neglected in
 433 monetary policy decisions, it leads to quantitative and qualitative policy mistakes. In
 434 order to further characterize these consequences, it is useful to focus on the *persistence*
 435 and *volatility* of y , i , and π and studying how they change as a function of sectoral
 436 price stickiness asymmetry.

437 Starting with persistence, this feature of the dynamics is here measured using, for
 438 each variable, the sum of the first five autocorrelation coefficients¹². Figure 5, Panel
 439 a, shows that the persistence of y , i , and in particular π , monotonically decreases for
 440 increasing degrees of sectoral price stickiness asymmetry.

441 [INSERT FIGURE 5 HERE]

442 Specifically, moving from $\alpha_m = \alpha_s = 0.7$ to $\alpha_m = 0.5$, $\alpha_s = 0.9$, the persistence
 443 of y , i and π decreases of 2.6%, 6.9%, and 24% respectively. Noting that in the two

¹²The autocorrelation coefficients have been computed in presence of the whole set of considered shocks.

sectors inflation persistence exhibits opposite behaviors, the stark fall in aggregate 444
inflation persistence depends on the major role played by the sector where prices are 445
more flexible, i.e. the m-sector. Indeed, the larger the asymmetry, the more the 446
m-sector is exposed to the destabilizing impact of the shock and to the stabilizing 447
impact of the switching demand mechanism; both factors contracting persistence. In 448
contrast, the larger the asymmetry, the less the s-sector is exposed to the destabi- 449
lizing impact of the shock and to the stabilizing impact of the switching demand 450
mechanism¹³. The decrease in inflation persistence leads, in turn, to a decrease in 451
the interest rate persistence via the Taylor rule. 452

The finding that heterogeneity in price stickiness leads to less inflation persistence 453
questions the ability of the standard New Keynesian Phillips curve in fitting the 454
macroeconomic evidence on inflation persistence, a view that this work shares with 455
Sheedy's (2007). 456

Let us now turn to the volatility of y , i , and π as a function of the asymmetry 457
in sectoral price stickiness. The impulse response analysis signalled that introduc- 458
ing sectoral asymmetry in price stickiness tends to reduce the deviations of these 459
variables from their long run values. Consistently, Figure 5, Panel b shows that 460
the unconditional standard deviations of these variables monotonically fall when the 461
asymmetry in sectoral price stickiness increases¹⁴. In particular, the variability of π 462
and i decreases of 38.2% leading to a decrease of the variability of y of the 31.3%. 463
This finding matters as indicates that sectoral asymmetry in price stickiness plays a 464
key role in determining the variability of the endogenous variables. 465

These results also imply that two economies that share the average but not the 466
variance of the sectoral price stickiness exhibit very different persistence and variabil- 467
ity of the endogenous variables. 468

At this point it is interesting to relate the current results to the inflation targeting 469
operating procedure in use at several central bank (Svensson 2010). Sketching this 470
procedure, first the staff computes alternative distribution forecasts associated with 471
different interest rate paths minimizing a standard loss function. Then the Board 472
selects the policy associated with the specific distribution forecast that suits best its 473
preferences. The current results thus suggest that one-sector models are not suitable 474
to obtain projections of the target variables since they can be highly misleading as 475

¹³This depends on the elasticities of sectoral inflations to real shocks and the relative price as explained above.

¹⁴In the Figure, the inflation and the interest rate volatility curve tend to overlap so that the inflation curve is hidden.

476 to their persistence and volatility.

477 Finally, these theoretical results are in line with the empirical evidence related
478 to French data found by Imbs et al. (2011). Indeed, they show that in presence of
479 sectoral heterogeneity in price stickiness, sectoral estimates imply policies associated
480 with volatilities in output gap and inflation that halve the ones implied by aggregate
481 data.

482 **4 Shocks and policy pitfalls of one-sector model** 483 **with price stickiness heterogeneity**

484 The role played by price stickiness heterogeneity in the transmission of monetary
485 shocks has been pioneered by Carvalho (2006). Therein the important result has been
486 to show that the presence of sectoral differences results in *larger* and *more persistent*
487 effects of monetary shocks on the output gap. Comparing the current findings with
488 Carvalho's we observe that they tend to be opposite: with real shocks the presence of
489 sectoral differences in prices stickiness results in *smaller* effects on the output gap¹⁵.

490 Two questions then naturally arise: why does this happen? and is there any gen-
491 eral result holding for *any* exogenous shock that can be useful for policy analysis? To
492 address the first question we simulate the response of the current model to a monetary
493 shock and illustrate the result using the previous AD-AS graphical framework.

494 Adding a monetary shock to the Taylor rule specified as $m_{t+1} = \gamma_m m_t + \varepsilon_{t+1}^m$ the
495 forward solution of the AD given by (19) is replaced by

$$\pi_t = - \left(\frac{1 + \tilde{\sigma}\delta_2}{\tilde{\sigma}\delta_1} \right) y_t + \frac{\delta_1 - 1 - \tilde{\sigma}\delta_2}{\tilde{\sigma}\delta_1} \sum_{\tau=1}^{\infty} \delta_1^{-\tau} y_{t+\tau|t} - \frac{1}{\delta_1} \sum_{\tau=0}^{\infty} \left(\frac{1}{\delta_1} \right)^{\tau} m_{t+\tau|t}. \quad (22)$$

496 Equation (22) shows that a monetary policy shock shifts the AD in the (π_t, y_t)
497 space. Since the position of the AD depends also on the expected path of the output
498 gap, if breaking the symmetry changes this path then there will be a different impact
499 of the monetary shock. To illustrate the relevant role played by the asymmetry
500 assumption, Figure 6 considers a contractionary monetary policy shock.

501 [INSERT FIGURE 6 HERE]

¹⁵Previous analysis has shown that output gap persistence falls too, but to a minor extent.

With both symmetry and asymmetry, the shock leads to a *negative* expected path for the output gap which, in turn, leads to a *negative* expected path for inflation. This leads the AD to shift left and the AS to shift down, point B. What happens when we break the symmetry? breaking the symmetry the AS shifts up and the AD shifts further left, point C (low asymmetry) or D (high asymmetry). Starting with the AS, the deflationary impact of the contractionary monetary policy will be smaller in the sector with stickier prices. Thus, the relative price q leaves the steady state and increases. This, in turn, exerts a positive effect on aggregate inflation contrasting the shift down of the AS determined by the negative expected output gap¹⁶.

Turning to the AD, why does it shift further left when we break the symmetry? Because, as explained above the switching demand mechanism prevents the expected inflation path from deviating from its steady state value as much as with symmetry. This implies that the policy response via the Taylor rule to inflation is smaller than under symmetry. As a result, the initial impact of the shock on the output gap is larger and the subsequent path to its steady state takes longer.

Let us now turn to the second question: is there any general result useful for policy analysis that we can apply to any exogenous shocks in presence of sectoral differences in price stickiness? Drawing on the AS-AD illustrations for the cost-push and monetary shock, Figures 4 and 6, we can note that, independently on the nature of the shock, breaking the symmetry contracts the outward shift of the AS. This contraction is due to the buffering device activated by the switching demand mechanism¹⁷. It is easy to show that this result is general in that applies to technology and preference shocks too. Thus, with any shock, sectoral differences in price stickiness lead to smaller and less persistent effects on inflation. This suggests three policy pitfalls for policy makers grounding policy decisions on one-sector model when the actual economy features sectoral differences in price stickiness. First, with *any* shock the one-sector model will *overvalue* the variation and persistence of inflation. Second, with real shocks the one-sector model will *overvalue* the variation of the output gap.

¹⁶Since the elasticity of current inflation to q is positive due to the fact that with asymmetry $\kappa_s < \kappa_m$, the q effect on aggregate inflation is positive.

¹⁷Formally, in both cases the shock leads to an expected path for q whose sign is opposite to the sign of the elasticity of π_t with respect to the shock at issue. Indeed, with a cost-push shock that has a direct positive impact on π_t it happens that q exhibits a negative path. With a monetary shock that has an indirect negative impact on π_t (via the expected path for the output gap), it happens that q exhibits a positive path.

530 Third, with monetary shocks the one-sector model will *undervalue* the variation and
531 persistence of the output gap. Table 3 reports these policy pitfalls.

532 **5 Concluding remarks**

533 Using a two-sector New-keynesian model, this paper investigates how sectoral het-
534 erogeneity in price stickiness affects the dynamics of the economy in presence of real
535 shocks.

536 When sectoral symmetry is broken, the relative price between sectoral goods ap-
537 pears into the working of economy and significantly alters its response to exogenous
538 shocks via the demand channel. As a result, asymmetry in sectoral price stickiness
539 on the one hand leads to an important fall in the persistence of aggregate inflation
540 and to a moderate fall in the persistence of the interest rate. On the other hand,
541 it leads to an important fall in the volatility of inflation, the interest rate and the
542 output gap. Thus, two economies differing in the dispersion of sectoral asymmetry
543 but not in the mean may exhibit very unlike volatility and persistence.

544 This finding has important monetary policy implications. Disregarding the dis-
545 persion in sectoral price stickiness leads policymakers first to overvalue the variation
546 and persistence of inflation; second, to overvalue the variation of the output gap. We
547 show that this wrong evaluation of the aggregate dynamics can lead policymakers
548 to relevant mistakes in presence of real shocks. They can follow over the first ten
549 quarters subsequent the shock a policy that is qualitatively different from the correct
550 one, or a policy that is up to 150% tighter or easier than the correct one. These
551 implications suggests that policy decisions should seriously account for the presence
552 of sectoral price stickiness. Moreover, they highlight the relevance of sectoral data
553 for economic estimations and forecasts.

554 Considering a country or monetary union where regions or states feature differ-
555 ent degree of price stickiness, the current findings question the appropriateness of
556 a monetary policy based on the average degree of price stickiness. Further analysis
557 will investigate this point as well as how sectoral size and strategic complementarities
558 affect the relation between sectoral asymmetry in price stickiness and the dynamics
559 of the economy.

560 **Acknowledgments**

561 I thank for comments Guido Ascari, Mustafa Caglayan, Alpay Filiztekin and

Kostas Mouratidis. I have also benefited from a useful discussion with Paul Levine 562
and Paul Mizen, and seminar participants at the University of Surrey, Nottingham 563
and Pavia. Any mistake is my responsibility. 564

References

- 565
566 Altissimo, F., Ehrmann, M., Smets, F., 2006. Inflation Persistence and Price-Setting
567 Behaviour in the EuroArea —A Summary of the IPN Evidence. Occasional Paper
568 Series 46, European Central Bank.
- 569 Aoki, K. 2001. Optimal monetary policy responses to relative-price changes. *Journal*
570 *of Monetary Economics* 48, 55-80.
- 571 Ascari, G., A. Flamini, and L. Rossi, 2012, Nominal Rigidities, Supply Shocks and
572 Economic Stability. DEM Working Papers Series 024, University of Pavia, Depart-
573 ment of Economics and Management.
- 574 Benigno P. 2004. Optimal monetary policy in a currency area. *Journal of Interna-*
575 *tional Economics* 63 (2004) 293– 320.
- 576 Bils, M. and P. Klenov (2002), "Some Evidence on the Importance of Sticky Prices",
577 *Journal of Political Economy* 112, 947-985.
- 578 Blinder, A., E. Canetti, D. Lebow and J. Rudd (1998), *Asking about Prices: A New*
579 *Approach to Understanding Price Stickiness*, Russell Sage Foundation.
- 580 Calvo G. A. 1983. Staggered prices in a utility-maximizing framework. *Journal of*
581 *Monetary Economics* 12(3): 383–398.
- 582 Carvalho C. 2006. "Heterogeneity in Price Stickiness and the Real Effects of Monetary
583 Shocks," *The B.E. Journal of Macroeconomics*, Berkeley Electronic Press, vol. 0(1),
584 pages 1.
- 585 Christiano, Lawrence J., Martin Eichenbaum, and Charles L. Evans. (2005) "Nom-
586 inal Rigidities and the Dynamic Effects of a Shock to Monetary Policy." *Journal of*
587 *Political Economy*, 113, 1–45.
- 588 Dhyne, E., L. Alvarez, H. Le Bihan, G. Veronese, D. Dias, J. Hoffman, N. Jonker, P.
589 Lunnemann, F. Rumler and J. Vilmunen (2006), "Price Changes in the Euro Area
590 and the United States: Some Facts from Individual Consumer Price Data," *Journal*
591 *of Economics Perspective* 20, 171-192.

- Dixon H. and E. Kara 2010a. Contract length heterogeneity and the persistence of monetary shocks in a dynamic generalized Taylor economy. forthcoming *European Economic Review*. 592-594
- Dixon H. and E. Kara 2010b. Can We Explain Inflation Persistence in a Way that Is Consistent with the Microevidence on Nominal Rigidity? *Journal of Money, Credit and Banking*, Vol. 42, No. 1, 151-170. 595-597
- Flamini A., G. Ascari, L. Rossi 2012. Industrial Transformation, Heterogeneity in Price Stickiness, and the Great Moderation. DEM Working Papers Series 025, University of Pavia, Department of Economics and Management. 598-600
- Fernández-Villaverde J., P. Guerrón-Quintana and J. Rubio-Ramírez 2010 "The New Macroeconometrics: A Bayesian Approach" in A. O'Hagan and M. West ,eds., *Handbook of Applied Bayesian Analysis* (2010), Oxford University Press. 601-603
- Imbs, J., Jondeau, E., Pelgrin, F. 2011. Sectoral Phillips curve and the aggregate Phillips curve. *Journal of Monetary Economics* 58 (4), pp. 328-344. 604-605
- Kara, E., 2010. "Optimal monetary policy in the generalized Taylor economy," *Journal of Economic Dynamics and Control*, Elsevier, vol. 34(10), pages 2023-2037. 606-607
- Nakamura, E., Steinsson, J., 2008. Five facts about prices: a reevaluation of menu cost models. *Quarterly Journal of Economics* 123, 1415–1464. 608-609
- Sheedy, K. D. (2007) Inflation Persistence when Price Stickiness Differs Between Industries. CEP Discussion Paper No 838. 610-611
- Smets F. and R. Wouters (2003). An Estimated Dynamic Stochastic General Equilibrium Model of the Euro Area. *Journal of the European Economic Association* 1 (5), 1123-1175. 612-614
- Smets, F., and R. Wouters. (2007). "Shocks and Frictions in US Business Cycles: A Bayesian DSGE Approach." *American Economic Review*, 97(3): 586–606. 615-616
- Svensson L. E. O. (2010), Inflation Targeting, Friedman, B. M., and M. Woodford, eds., *Handbook of Monetary Economics*, Volume 3a and 3b, North-Holland. 617-618

619 Woodford, M. (2011). "Optimal Monetary Stabilization Policy" in Benjamin M.
620 Friedman & Michael Woodford (ed.), Handbook of Monetary Economics, Elsevier,
621 edition 1, volume 3, number 3.

TABLE 1 Alternative calibrations

Panel a -benchmark				Panel b -robustness	
η	0.7	n_s	0.5	η	0
$\tilde{\sigma}$	2/3	n_m	0.5	$\tilde{\sigma}$	1
β	0.9975	γ_a	0.95		
ρ	1	γ_ψ	0.95	ρ	8
ν	1.17	γ_c	0.95	ν	1
$1/\phi$	0.75	σ_ε^2	0.009 ²	$1/\phi$	1
θ	7.88			θ	8
δ_0	0.8				
δ_1	1.5				
δ_2	0.5/4				

622

TABLE 2. Average Policy Error for the First Ten Periods

		Shock		
623	Calibration	ε_ψ	ε_a	ε_c
	1	50%	50%	60%
	2	100%	100%	180%

TABLE 3. Policy pitfalls using one-sector models in presence of price stickiness asymmetry

		Shock	
		Real	Monetary
π	Overvaluation of persistence and variation		Overvaluation of persistence and variation
y	Overvaluation of variation		Undervaluation of persistence and variation

Figure 1. IRFs to real shocks in presence of symmetry and asymmetry in sectoral price stickiness

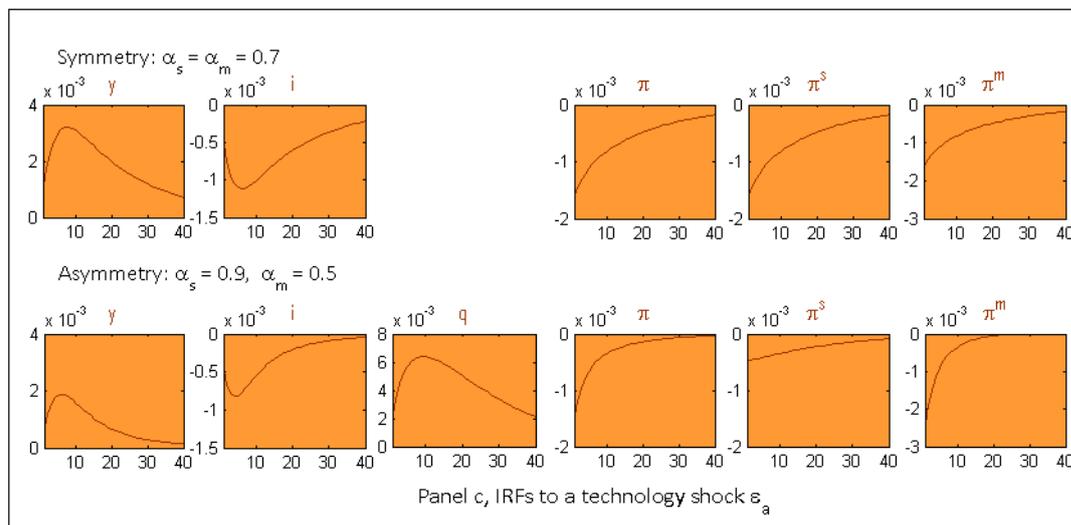
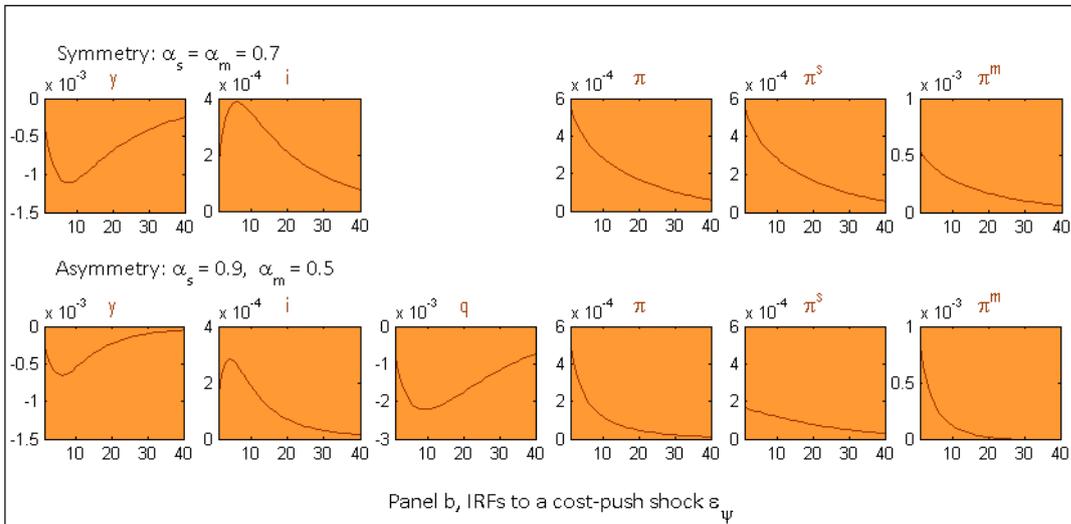
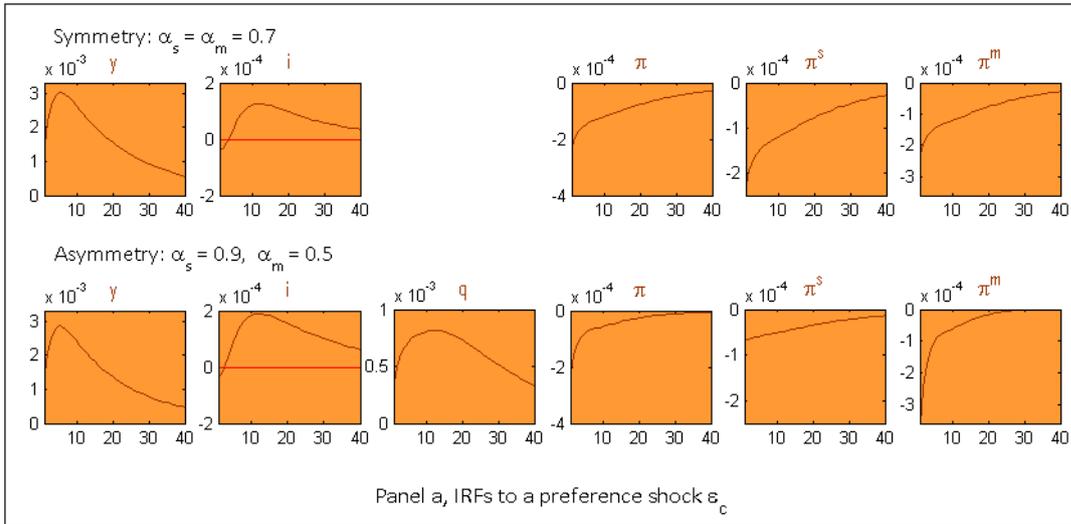


Figure 2. IRFs of the interest rate to real shocks under sectoral price stickiness symmetry (actual policy) and asymmetry (correct policy)

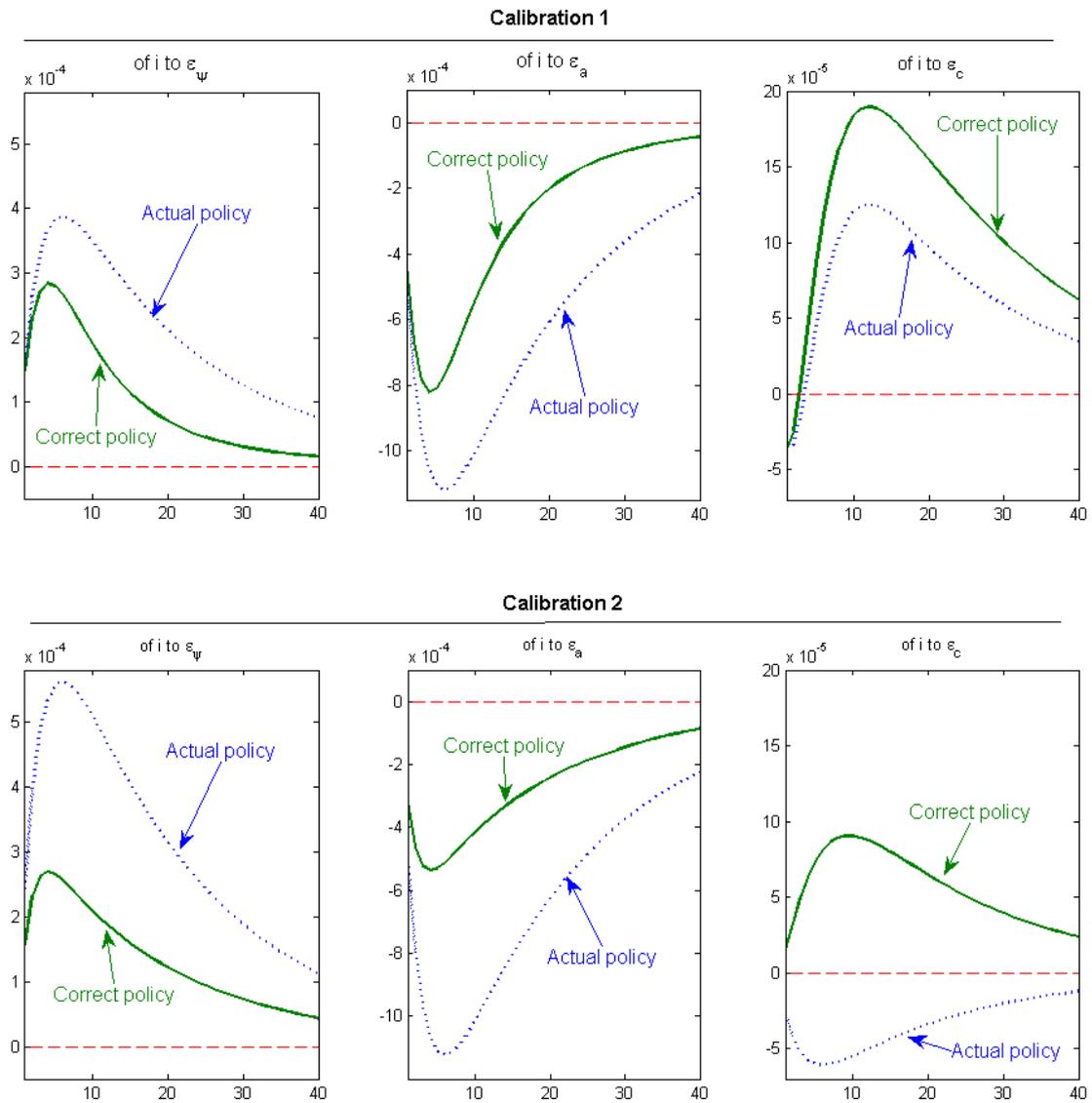


Figure 3. Monetary policy mistakes in presence of real shocks: percentage deviation of the actual policy from the correct policy

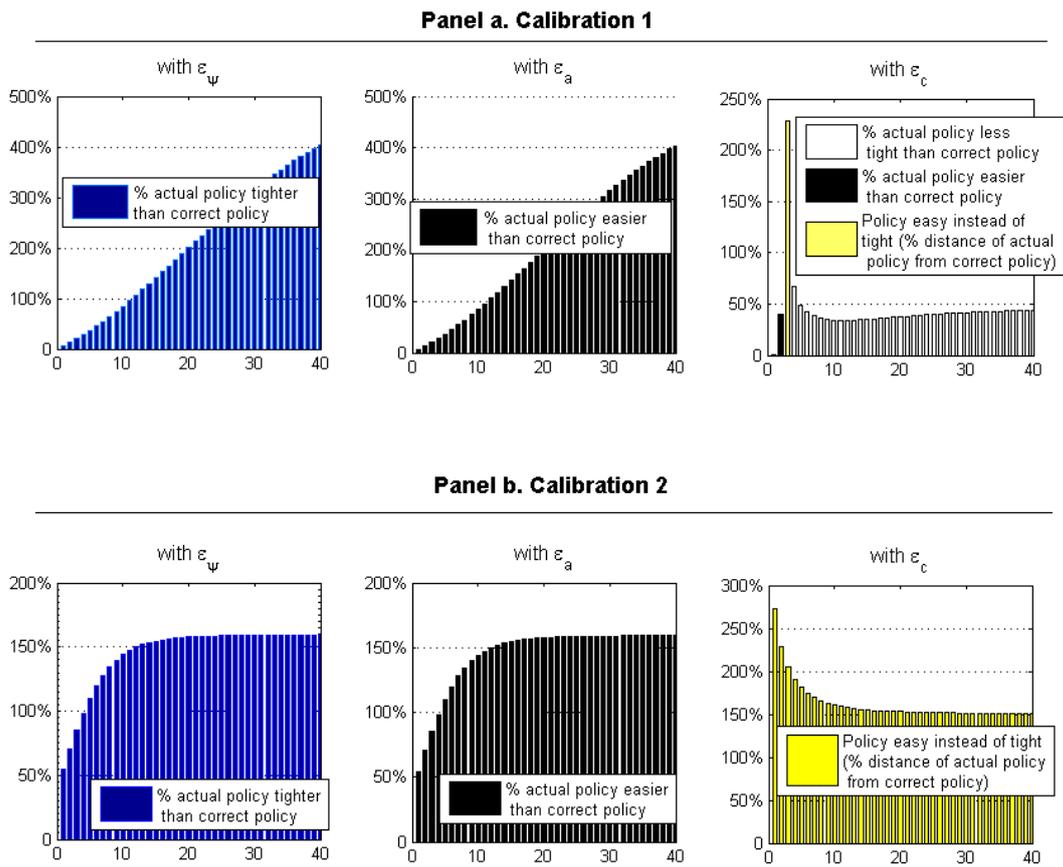


Figure 4. Steady state equilibrium and off-steady state equilibria after a cost-push shock with symmetry and asymmetry in sectoral price stickiness.

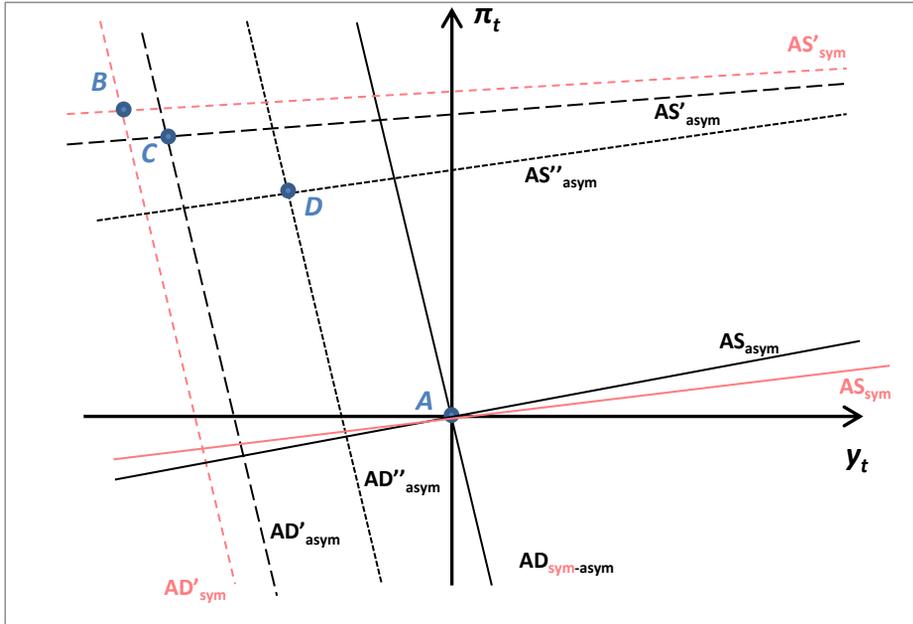
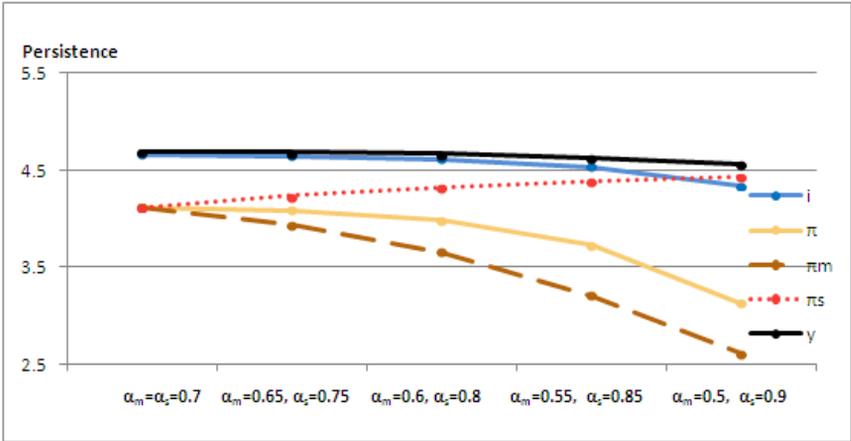
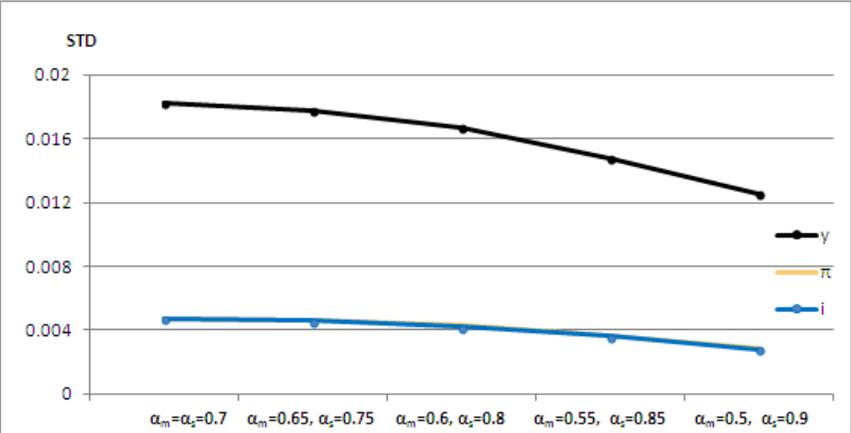


Figure 5. Persistence and volatility of the macrovariables vs asymmetry in sectoral price stickiness



Panel a. Persistence vs asymmetry in sectoral price stickiness



Panel b. Volatility vs asymmetry in sectoral price stickiness

Figure 6. Steady state equilibrium and off-steady state equilibria after a monetary shock with symmetry and asymmetry in sectoral price stickiness.

